

2.2 Randomization of a function on \mathbb{T}^d (and \mathbb{R}^d)

$$f(x) = \sum_{n \in \mathbb{Z}^d} \hat{f}(n) e^{inx}, \quad x \in \mathbb{T}^d.$$

- Let $\{g_n\}_{n \in \mathbb{Z}^d}$ be a collection of independent random variables with distributions of real and imaginary parts $M_n^{(r)}, M_n^{(i)}$ mean - zero
assume indep.

$$\int e^{r \cdot x} d\mu_n^{(r)}(x) \leq e^{cr^2}$$

$$\int e^{r \cdot x} d\mu_n^{(i)}(x) \leq e^{cr^2}, \quad \forall n \in \mathbb{Z}^d, r \in \mathbb{R}$$

\Leftrightarrow satisfied by standard \mathbb{C} -valued Gaussian r.v.'s,
Bernoulli (or uniform distri on $S^1 \cap \mathbb{C}$).

Rmk: In many situations, it suffices to assume (uniform) boundedness of the k^{th} moment $\mathbb{E}[|g_n|^k] \leq c < \infty$ ②

\Rightarrow Randomization f^ω of f :

$$f^\omega(x) = \sum_{n \in \mathbb{Z}^d} \underline{g_n(\omega)} \hat{f}(n) e^{inx}$$

Rmk: If f is real, i.e. $\hat{f}(t+n) = \overline{\hat{f}(n)}$, $\forall n \in \mathbb{Z}^d$.

$$\text{Let } I = \bigcup_{k=0}^{d-1} \mathbb{Z}^k \times \mathbb{Z}_+ \times \{0\}^{d-k-1} = " \mathbb{Z}^d / 2 "$$

$$\Rightarrow \mathbb{Z}^d = I \cup \{-I\} \cup \{0\}$$



Then, introduce $\{g_n\}_{n \in I \cup \{0\}}$

and set $g_{-n} = \overline{g_n}$, $\forall n \in \mathbb{Z}^d$. (g_0 is real-valued.)

$\Rightarrow f^\omega(x) = \sum_{n \in \mathbb{Z}^d} g_n(\omega) \hat{f}(n) e^{inx}$ is real-valued.

Moral: If $f \in H^s \setminus H^{s+\epsilon}$, then $f^\omega \in H^s \setminus H^{s+\epsilon}$ a.s. ③

i.e. there is no gain in differentiability.

• If $f \in L^2$, then $f^\omega \in L^p$, $\forall 2 \leq p < \infty$, a.s.

⇒ gain of integrability.

Lemma 2.1: (i) Given $f \in H^s(\mathbb{T}^d)$, let f^ω be the randomization defined above. Then, the following tail estimate holds:

$$P\left(\|f^\omega\|_{H^s} > \lambda\right) \leq C \exp\left(-c \frac{\lambda^2}{\|f\|_{H^s}^2}\right) \xrightarrow[\text{as } \lambda \rightarrow \infty]{} 0.$$

for all $\lambda > 0$. In particular, $f^\omega \in H^s$, a.s.

(ii) Let $f \in L^2(\mathbb{T}^d)$. Then, for any $\sqrt{p} \geq 2$, we have

$$P\left(\|f^\omega\|_{L^p} > \lambda\right) \leq C \exp\left(-c \frac{\lambda^2}{\|f\|_{L^2}^2}\right).$$

for all $\lambda > 0$. In particular, $f^\omega \in L^p$, a.s.

Paley-Zygmund '30

Kahane's book

Lemma 2.2: Let $\{g_n\}_{n \in \mathbb{Z}^d}$ be as above. Then, we have ④

$$\left\| \sum_{n \in \mathbb{Z}^d} c_n g_n \right\|_{L^p(\Omega)} \lesssim \sqrt{p} \|c_n\|_{\ell_n^2}.$$

Pf: We first prove

$$P(|\sum c_n g_n| > \lambda) \leq C e^{-c \frac{\lambda^2}{\|c_n\|_{\ell_n^2}^2}}.$$

$$\begin{aligned} \cdot \int e^{t \sum c_n g_n} dP &\stackrel{\text{indep}}{=} \prod_n \int e^{t c_n g_n} dP \\ &\leq \prod_n e^{\alpha (t c_n)^2} = e^{\alpha^2 t^2 \sum c_n^2}. \\ \Rightarrow P(\sum c_n g_n > \lambda) &\leq \frac{\mathbb{E}[e^{t \sum c_n g_n}]}{e^{t \lambda}} \leq \frac{e^{\alpha^2 t^2 \sum c_n^2}}{e^{t \lambda}} \\ \text{choose } t = \frac{\lambda}{2\alpha^2 \sum c_n^2} &= e^{-\frac{\lambda^2}{4\alpha^2 \sum c_n^2}}, \end{aligned}$$

$$\text{Similarly, } P(\sum c_n g_n < -\lambda) \leq C e^{-c \frac{\lambda^2}{\|c_n\|_{\ell_n^2}^2}}.$$

$$\Rightarrow \left\| \sum c_n g_n \right\|_{L^p(\mathbb{R})}^p \leq p \int_0^\infty \lambda^{p-1} e^{-c \frac{\lambda^2}{\|c_n\|_{\ell^2_n}^2}} d\lambda.$$

(5)

Ch. of var
 $\lesssim p \left(\sum c_n^2 \right)^{p/2} \underbrace{\int_0^\infty \lambda^{p-1} e^{-\lambda^2/2} d\lambda}_{\sim (p-1)!!} \sim (p-1)!! = \frac{(2p)!}{2^p \cdot p!}$

\square

Aside $\int_R x^k e^{-x^2/2} dx$, k even.

\Leftarrow IBP does not help. Use the moment generating function

$$\mathbb{E}[e^{tx}] \stackrel{\text{Ch. f var}}{=} e^{\frac{t^2}{2} + t^2} = \sum_{n=0}^{\infty} \frac{1}{n! 2^n} t^{2n}$$

$$\mathbb{E}\left[\sum_{k=0}^{\infty} \frac{(tx)^k}{k!}\right] = \sum_{k=0}^{\infty} \frac{\mathbb{E}(x^k)}{k!} t^k. \text{ Compare the coefficients.}$$

pf of Lemma 2.1 : By Minkowski's integral inequality. ⑥

$$\begin{aligned} \# P \geq 2 \quad & \| \| f^\omega \|_{H^s} \|_{L^p(\Omega)} \leq \| \| \sum_n \underbrace{\langle m \rangle^s \hat{f}(m) e^{inx}}_{c_n(x)} g_m \|_{L^p(\Omega)} \|_{L_x^2} \\ & \| \langle v \rangle^s f^\omega \|_{L_x^2} \stackrel{\text{Lem 2.2}}{\lesssim} \sqrt{p} \left\| \underbrace{\| \langle m \rangle^s \hat{f}(m) \|_{l_n^2}}_{= \| f \|_{H^s}} \right\|_{L_x^2(\mathbb{T}^d)} \\ & \sim \sqrt{p} \| f \|_{H^s}. \end{aligned}$$

By Chebyshev's Ineq,

$$P(\| f^\omega \|_{H^s} > \lambda) < \left(\frac{C_0 P^{1/2} \| f \|_{H^s}}{\lambda} \right)^p, \quad \# P \geq 2.$$

Let $P = \left(\frac{\lambda}{C_0 e \| f \|_{H^s}} \right)^2$. If $P \geq 2$, then we have

$$P(\| f^\omega \|_{H^s} > \lambda) < e^{-P} = e^{-C \frac{\lambda^2}{\| f \|_{H^s}^2}}$$

(7)

If $p = \left(\frac{\lambda}{C_0 e \|f\|_{H^s}} \right)^2 \leq 2$, then

choose C s.t. $C e^{-2} \geq 1$. Then, we have

$$P(\|f^\omega\|_{H^s} > \lambda) \leq 1 \leq C e^{-2} \leq C e^{-c \lambda^2 / \|f\|_{H^s}^2}$$

(Alternative way: Establish $\mathbb{E}[e^{\alpha \|f^\omega\|_{H^s}^2}] < \infty$

by Taylor expansion + $\|\|f^\omega\|_{H^s}\|_{L^p(\Omega)} \lesssim \sqrt{p} \|f\|_{H^s}$

Given $p \geq 2$, $\|\|f^\omega\|_{L_x^p}\|_{L_x^q(\Omega)} \stackrel{(ii)}{\leq} \|\sum_n \underbrace{\hat{f}(n) e^{inx}}_{c_n} \cdot g_n\|_{L_x^q(\Omega)} \|_{L_x^p}$
let $q \geq p$.

$$\begin{aligned} &\lesssim \sqrt{q} \|\underbrace{\hat{f}(n)}_{d_n}\|_{L_x^2} \|_{L_x^p} \\ &= \|f\|_{L_x^2} \\ &\sim \sqrt{p} \|f\|_{L_x^2}. \quad \# q \geq p. \end{aligned}$$



Lemma 2.3 : Suppose $f = \sum \hat{f}(m) e^{inx} \in H^s \setminus H^{s+\varepsilon}$ for some $\varepsilon > 0$. ⑧

Suppose $\exists c > 0$ s.t.

$$\limsup_{|n| \rightarrow \infty} P(\{|g_n| \leq c\}) \leq 1 - \delta < 1$$

(e.g. satisfied if g_n is i.i.d. and $g_n \neq 0$)

$$\text{Then, } P(f^\omega \in H^{s+\varepsilon}) = 0$$

Pf : For simplicity, assume g_n is real-valued.

$$\begin{aligned} \int e^{-\|f^\omega\|_{H^{s+\varepsilon}}^2} dP &\stackrel{\text{indep}}{=} \prod_m \left(\int e^{-c^2 n^{2(s+\varepsilon)} |\hat{f}(m)|^2 |g_n|^2} d\mu_n \right) \\ &\leq \prod_n \left(\mu_n(-c, c) + \underbrace{e^{-c^2 n^{2(s+\varepsilon)} |\hat{f}(m)|^2}}_{=: \alpha_n} (1 - \mu_n(-c, c)) \right) \\ &= \prod_n (\mu_n(-c, c)(1 - \alpha_n) + \alpha_n) \\ &\leq \prod_n ((1 - \delta)(1 - \alpha_n) + \alpha_n) = \prod_n (1 - \delta(1 - \alpha_n)) \stackrel{\text{**}}{=} 0 \end{aligned}$$

$$\Rightarrow \|f^\omega\|_{H^{s+\varepsilon}} = \infty, \text{ a.s.}$$

(9)

Proof of $\star\star$: By assumption, $\sum n^{2(s+\varepsilon)} |\widehat{f}(n)|^2 = \infty$.

Then, we claim $\sum (1-\alpha_n) = \infty$

$$\sum (1 - e^{-c^2 n^{2(s+\varepsilon)} |\widehat{f}(n)|^2})$$

- If $n^{2(s+\varepsilon)} |\widehat{f}(n)|^2 \rightarrow 0$ as $|n| \rightarrow \infty$, then $\sum (1 - \alpha_n) \geq \sum (1 - \varepsilon) = \infty$.
 - Otherwise, $\exists N$ s.t. $c^2 n^{2(s+\varepsilon)} |\widehat{f}(n)|^2 \leq 1$, $\forall |n| \geq N$
- $$\Rightarrow \sum_{|n| \geq N} (1 - \alpha_n) \leq C \sum_{|n| \geq N} n^{2(s+\varepsilon)} |\widehat{f}(n)|^2 = \infty.$$
- (b/c $e^{-x} \leq 1 - Cx$ for $0 \leq x \leq 1$. (for some $C \ll 1$))

$$\begin{aligned} \text{Then, } \left(\prod_n (1 - \delta(1 - \alpha_n)) \right)^{-1} &= \prod_n \left(1 + \frac{\delta(1 - \alpha_n)}{1 - \delta(1 - \alpha_n)} \right) \\ &\leq \prod_n \left(1 + 2\delta(1 - \alpha_n) \right) \quad \text{by choosing } \delta \leq \frac{1}{2}. \\ &= \infty. \end{aligned}$$

$$\left(\prod_n (1 + a_n) < \infty \Leftrightarrow \sum a_n < \infty. \right)$$